



Federal Reserve Bank of Chicago

**Immigrant-Native Differences in  
Financial Market Participation**

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## **Immigrant-Native Differences in Financial Market Participation**

The goal of this paper is to investigate the prospects for the wealth assimilation of immigrants by studying the financial market behavior of U.S. immigrants, compared to the native-born. Compared to similar natives, immigrants are less likely to own a wide range of financial assets, including savings and checking accounts. Immigrant status also has a significant impact on transitions out of account ownership. We find that lower rates of financial market participation tend to persist even for immigrants who have lived in the U.S. for several years. Our results suggest that a large share of the immigrant-native gap in financial market participation is driven by group differences in education, income, and geographic location. For a given immigrant, the likelihood of financial market participation decreases with higher levels of ethnic concentration in the metropolitan area.

Keywords: Immigrants, financial markets, ethnic concentration, location

## **I. Introduction**

A central question facing researchers and policymakers is the extent to which immigrants will adapt to economic, social, and political life in United States. One crucial facet of economic and social well-being is wealth. While there is a rich literature that examines the sources of immigrant-native differences in labor market, health, and educational outcomes, relatively little is known about the determinants of wealth differences between immigrants and the native born. In a growing number of studies, researchers have documented that immigrants have substantially lower wealth levels and that differential patterns in asset holdings can explain a large share of the immigrant-native wealth gap (Amuedo-Dorantes and Pozo, 2002; Hao, 2004, Cobb-Clark and Hildebrand, 2006a, 2006b; Krivo and Kaufman, 2004). In particular, the median wealth levels of natives are estimated to be about 2.3 times higher that of immigrants, and immigrants are less likely to hold financial and real estate wealth compared to natives (Cobb-Clark and Hildbrand, 2006a).

In this paper, we investigate the factors that influence the decision of immigrants and natives to hold wealth in a particular form—financial assets. By focusing on the extensive margin, rather than the intensive margin, we hope to shed light on the reasons for the wide disparities in immigrant- native wealth holdings. Based on recent data, financial wealth holdings are the second largest component of household wealth in the U.S., after housing, and account for 42% of overall household assets in 2001 (Survey of Consumer Finances, 2001). The relative liquidity of financial assets means that they can play an important role in allowing households to self-insure against negative income shocks induced, for example, by job loss, illness, and marital disruption. Beyond their role in helping households cope with income uncertainty, financial assets also tend to be associated with high expected returns over time and can contribute significantly to understanding long-run differences in wealth accumulation. Several empirical studies have shown that the ownership and the value of financial assets is correlated with a wide range of economic and social decisions, including investments in health (Roberts and House, 1996), well-being at retirement (Bender, 2004), home ownership (Haurin et al., 1997), business formation (Holtz-Eakin et al., 1994), children’s educational outcomes (Duncan and Brooks-Gunn, 1997; Mayer, 1997), and political participation (Scanlon and Page-Adams, 2001).

We use panel data from the 1996 – 2000 Survey on Income and Program Participation (SIPP) to analyze financial market decisions of immigrants relative to those of natives and to

estimate the impact of being an immigrant on the likelihood of transitions into and out of financial market participation. In addition to documenting differences in immigrant versus native financial market participation, we show how duration of stay in the United States impacts immigrant behavior relative to the native-born. Based on unconditional means, immigrants are much less likely to hold U.S. financial assets compared to their native-born counterparts. For example, 40% of immigrants have a savings account, compared to 55% of the native-born. There is a similar gap in the percentage of immigrants and native-born who hold an interest-bearing checking account: 22% versus 36%. While 12% of immigrants own stock, ownership rates are much higher among the native-born at 25%. We find that immigrants also report lower rates of ownership of a wide range of additional financial assets, including individual retirement accounts and mutual funds when compared to the native-born.

Controlling for education, income, and other individual and household characteristics, the financial market behavior of immigrants remains significantly different from that of similar native-born individuals. An important advantage of the SIPP panel is that we observe not just financial asset holdings at one point in time, but also transitions into and out-of financial market participation at frequent intervals (every four months) for both savings and interest-bearing checking accounts. Taken together, we find that immigrants transition out of financial asset ownership at higher rates than similar natives and that recent immigrants are significantly less likely to enter into savings and checking account ownership compared to natives. We present additional evidence that the explanation for differential behavior of immigrants relative to natives has to do with the characteristics of the geographic area where a given immigrant resides. For a given immigrant, the likelihood of financial market participation decreases with the percentage of the population in a given metropolitan area from the same origin country as the immigrant in question. One potential explanation for these results is that social interactions may play an important role in determining whether immigrants participate in financial markets or that immigrant networks provide informal substitutes for formal financial markets. Like wage growth (Borjas, 1998 and 2000), human capital accumulation, and language proficiency (Chiswick and Miller, 1996), immigrant participation in formal financial markets appears to be inhibited when there is a large network of immigrants from the same country of origin to interact with.

The rest of the paper is organized as follows. Section II describes the conceptual framework and empirical methods. We describe the SIPP data and summarize the data on the financial market participation of immigrants relative to natives in Section III. In Section IV, we present results. Section V presents conclusions.

## **II. Understanding Immigrant-Native Differences in Financial Market Participation**

A simple life-cycle model is a useful starting point for exploring immigrant-native differences in financial market behavior (Modigliani and Brumberg, 1954). Individuals, whether immigrant or native-born select a portfolio from a wide range of assets, comparing returns, transaction costs, risk profiles, and liquidity in order to smooth consumption over time.<sup>1</sup> This approach suggests that differences in financial market participation between immigrants and the native-born may be driven (at least in part) by differences in household income, education, age, and family structure. The decision to hold a particular asset will also depend on information, tastes and preferences, and the degree of risk aversion, which may differ across otherwise similar immigrants and natives. We recognize that it may be important to account for additional sources of immigrant-native differences, including race and ethnicity, legal status, English language skills, years of U.S. experience, and patterns of residential settlement, which are likely to affect financial decisions (Amuedo-Dorantes et al., 2005; Cobb-Clark and Hildenbrand, 2006a). In the section below, we describe the empirical models used to estimate the gap between immigrant and native-born financial market participation.

### **Empirical Specification**

#### ***1. Immigrant-Native Differences in Financial Market Participation***

The basic specification investigates the likelihood that an individual holds a given financial asset during a given period. The benefits and costs associated with the financial market participation for individual  $i$  living in a destination community  $j$  at time  $t$  can be defined as  $U_{ijt}$ , which is a function of  $(Z_{ijt})$ , a vector of socio-economic and demographic variables including, education, race, income, household size, and other control variables. In addition, for immigrants,  $U_{ijt}$  may be a function of immigrant status,  $(I_i)$ , and duration of stay in the United States,  $D_i$ . The net benefits and costs associated with holding a given financial asset may also vary by community,

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<sup>1</sup> There are additional motives for savings and wealth accumulation, which include precautionary motives and the desire to leave bequests.

$C_j$ , with time  $v_t$ , and be subject to an error term,  $\varepsilon_{ijt}$  that is particular to the individual. For each time period,  $U_{ijt}$  can be measured as:

$$U_{ijt} = \alpha + \beta_1 Z_{ijt} + \beta_2 I_i + \beta_3 (D_i * I_i) + \gamma_j * C_j + v_t + \varepsilon_{ijt} \quad (1)$$

We do not observe  $U_{ijt}$ , but we observe whether the household has participated in a given financial market. Thus, we observe:

$$P_{ijt} = 1 \text{ if } U_{ijt} > 0 \\ = 0 \text{ otherwise}$$

Equation (1) represents the fully specified model. We build up to this model and first estimate a parsimonious specification, which includes individual characteristics and an indicator variable for immigrant status. We use a maximum likelihood logit model to estimate the probability that an individual has participated in a given financial market in the survey period. The parameter on the immigrant indicator,  $\beta_2$ , will capture the effect of being an immigrant on the likelihood of holding a given financial asset, after having controlled for time in the U.S. and socio-economic and demographic characteristics. The parameter on the cohort of arrival variable,  $\beta_3$ , measures how duration of stay in the U.S. affects the immigrant's likelihood of participating in a given financial market. The set of parameters,  $\gamma_j$ , measure community level fixed-effects. We identify the community as the Metropolitan Statistical Area (MSA) where an individual resides. In our empirical estimates, we include MSA-level fixed-effects, which capture the effect of community variables such as the density of formal financial institutions in the MSA, employment conditions, and other economic attributes of the MSA. We also include time controls in all estimates to capture any time variation in financial market participation over the sample period. All reported standard errors are adjusted to allow for correlation across observations for a given individual.

## 2. *Immigrant-Native Differences in Exit and Entry in Financial Markets*

Financial market participation for both immigrants and natives, measured at a given point in time will depend on both exit and entry rates into holding a given financial asset. In the empirical analysis, we also focus on the impact of immigrant status on transitions into and out of financial market participation as these estimates can provide insights into why immigrant financial behavior differs from that of natives. Let  $V_{ijt}$  represent the net benefits of entry (or exit) into the use of a given financial service from time  $t - 1$  to time  $t$ . The net benefits of entry (or exit) are defined to be a function of individual and household characteristics at time  $t - 1$ ,

immigrant status, year of arrival controls, as well as community controls. Specifically, we measure  $V_{ijt}$  as:

$$V_{ijt} = \alpha + \beta_1 Z_{ijt-1} + \beta_2 I_i + \beta_3 (D_i * I_i) + \gamma_j * C_j + v_t + \eta_{ijt} \quad (3)$$

We do not observe the net benefits of entry or exit; instead we know whether the household has experienced a transition into (or out of) the use of a given financial service. Hence, we estimate using logit maximum likelihood:

$$\begin{aligned} E_{ijt} &= 1 \text{ if } V_{ijt} > 0 \\ &= 0 \text{ otherwise,} \end{aligned} \quad (4)$$

For estimates of entry, the dependent variable is equal to one if the individual reports owning an account at time  $t$  and not owning an account at time  $t - 1$ . The dependent variable is equal to zero if the individual reports no account ownership at time  $t$  and at time  $t - 1$ . For exit, the dependent variable is equal to one if an individual reports ownership at  $t - 1$  and no ownership at time  $t$  and is equal to zero if the individual reports ownership at both  $t$  and  $t - 1$ .<sup>2</sup> All of the transition estimates include the explanatory variables described in the discussion of the baseline results and standard errors are adjusted to allow for correlation across observations at the individual level.

### III. Data and Characteristics of Immigrants and the Native-born

The empirical analysis uses longitudinal data from the 1996-2000 waves of the Survey of Income and Program Participation (SIPP). The SIPP is a panel survey which provides detailed information about adults residing within households, and is conducted by the U.S. Census Bureau. The SIPP collects data by interviewing individual respondents (about 65,000 individuals) once a quarter about their economic experiences, including ownership of savings accounts, checking accounts, and stocks. The 1996 SIPP panel consists of twelve waves of interview questions, where the interview questions depend on the wave. We include only individuals who are 18 or older in our study. The analysis deals with individuals who reside in a Metropolitan Statistical Area (MSA). This allows us to control for MSA-level variation in the availability of financial services by including MSA controls in estimates of financial market

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<sup>2</sup> We should note that the entry estimates are restricted to those who report no ownership at time  $t - 1$  and that the exit estimates are restricted to those who do own an account at time  $t - 1$ .



behavior.<sup>3</sup> The sample includes, about 28,633 native-born, and 4,450 immigrants. Because we observe individuals multiple times, the total sample is made up of 356,769 quarterly observations.

In addition to information on financial asset holdings, the SIPP data include information on immigrant status, country of origin, and year of arrival in the U.S., coded into 5-year intervals to protect respondent confidentiality. The SIPP data are well-suited for this study because they include information on financial market behavior and on immigration. Other data sources available from the Bureau of the Census, or from the monthly Current Population Survey, contain a large number of immigrants and provide detailed information on immigration, but include very limited information on participation in financial markets or transitions into and out of ownership.<sup>4</sup> The immigrant population in the 1996 SIPP closely mirrors 2000 Census data on U.S. immigrants. Out of a total sample of 29,731 MSA residents, 14% are immigrants.<sup>5</sup>

Our analysis is conducted at the individual level to capture the extent to which immigrants and the native-born differ in the distribution of financial asset holdings.<sup>6</sup> The SIPP data provides detailed information for all adults in the household are interviewed on demographic characteristics, ownership of interest or dividend-earning financial accounts, and income. These data are available for each of the 12 waves, at approximately 3-month intervals. While the SIPP panel is relatively short, the large sample sizes available provide an opportunity to observe within-sample changes in financial asset holdings for both immigrants and the native-born.<sup>7</sup>

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<sup>3</sup> By focusing on an urban sample, we can also eliminate an important source of heterogeneity between immigrants and natives since about 75% of the SIPP immigrant sample lives in a MSA compared to about half of natives.

<sup>4</sup> The SIPP data do not include any information on remittances or the use of informal financial institutions. This makes it difficult to directly assess how participation in formal financial markets in the U.S. is impacted by immigrant financial ties to origin countries and the use and availability of informal financial substitutes.

<sup>5</sup> In the 2000 Census, 11.4% of the total population was born abroad. The higher percentage of immigrants that we find in our sample is due to the fact that we restrict our attention to MSA residents, and immigrants are more likely to live in metropolitan areas than in rural areas.

<sup>6</sup> An advantage of our approach is that we do not have to impose social norms about the degree to which assets are jointly held within the household as these may vary by country of origin. However, we find comparable results on immigrant-native differences in financial market behavior when we restrict the sample to household heads. These results are available upon request.

<sup>7</sup> With any data that tracks individuals and/or households over time, the problem of individuals dropping out of the sample during the course of data collection arises. Our analysis indicates that while immigrants drop out of the sample at higher rates between Wave 1 and Wave 2, after that patterns of attrition are fairly similar for natives and immigrants.

## A. Financial Market Participation

The empirical work focuses on two indicators of financial market participation: ownership of savings and interest-bearing checking accounts.<sup>8</sup> We emphasize savings and checking account ownership because these represent entry-level financial assets with relatively low barriers to participation. However, we also examine additional indicators of financial market participation including individual retirement accounts (IRA) or Keogh and mutual fund ownership, although mean ownership rates tend to be lower for these indicators of financial market participation. Figure 1 summarizes this information. The SIPP data provide information on whether a given survey respondent held a particular financial asset in the previous month at four month intervals for the duration of the panel. To allow a comparison across the two of the major components of asset holdings, we present immigrant-native gaps in homeownership (see Borjas, 2002; Krivo and Kauffman, 2004).

Table 1 summarizes patterns of financial asset ownership for immigrants and the native-born. Compared to the native-born, immigrants are less likely to hold mainstream financial assets. We find that ownership of savings accounts appears relatively widespread in the SIPP data, with 53% of the pooled immigrant-native-born sample reporting ownership of a savings account. However, only 40% of immigrants own a savings account compared to 55% of natives. Ownership of an interest-bearing checking account is less common, with only 34% of the sample reporting ownership. For checking accounts, the gap between immigrants and natives is even larger, with immigrant ownership rates of 22% being only 60% that of the native-born at 36%. While 18% of the pooled sample report stock ownership, only 9% of immigrants own stock compared to the ownership rate of the native-born at 20%.

Table 1 also reports summary statistics on exit and entry from the SIPP for savings, checking accounts, and stock. Transitions into and out of account ownership differ in important ways by immigrant-native status. Over the course of the panel, immigrants are less likely to participate in mainstream financial markets. For example, about 43% of immigrants report *never* owning a savings account throughout all 12 waves, compared to 30% of the native-born. We also note that for immigrants the percentage of immigrants who *never* owned a checking account at any time during the panel is about 66%, compared to 51% for the native-born. Similarly, for

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<sup>8</sup> We focus our attention on interest-bearing checking accounts because information on ownership of non-interest bearing checking accounts is available less frequently in the SIPP (approximately once a year). We obtain similar results when we combine both interest and non-interest bearing checking accounts (see Figure 1).

immigrants, the percentage of immigrants who *never* owned stock at any time during the panel is about 85%, compared to 71% for the native-born.

Immigrants report more volatility in their financial market participation status. We find that immigrants are more likely than the native-born to report exits from financial asset holdings. Specifically, exit rates out of savings and checking account ownership for immigrants are about 60% higher for immigrants than for the native-born. Immigrants also have lower rates of entry into savings and checking account, and stock ownership compared to the native-born.

## **B. Characteristics of Immigrants and the Native-born**

### **Socioeconomic and Demographic Characteristics**

Table 2 provides a detailed comparison of the characteristics of immigrants and the native-born. Compared to the native-born, immigrants are younger, more likely to be married, have more children, and more likely to be unemployed or economically inactive. Immigrants also tend to be less educated than the native born. Nearly 36% of the immigrant sample has not completed high school compared to only 15% of the native-born sample. However, the percentage of immigrants and the native-born who have an advanced degree is comparable at about 7%. Monthly per capita household income is significantly lower for immigrants compared to the native-born. For immigrants, average monthly per capita household income is \$1,619, compared to \$2,195 for the native-born. We also note that immigrants are more likely to be non-white. About 75% of the immigrant sample is non-white compared to about 23% of the native-born sample. Nearly 30% of the immigrant sample was born in Central America, while about 15% of the immigrant sample is of European descent. A sizeable share of the immigrants in the SIPP can be classified as recent immigrants, with almost 40% of the immigrants arriving in the U.S. after 1990.

## **IV. Results**

### **A. Baseline Findings for Participation in Financial Markets**

We present the baseline findings in Tables 3. The key dependent variables in the analysis are indicator variables that capture whether or not an individual owned a savings account (column 1) or a checking account (column 2) during the survey reference period. All of our estimates include MSA fixed-effects, as well as the following explanatory variables: age, age

squared, labor force status, per capita income, per capita income squared, marital status, the number of children in the household, sex, race, and education.<sup>9</sup> We present additional results for stock, individual retirement accounts and other financial assets in Appendix Table.<sup>10</sup>

### *Individual and Household Variables*

First, we discuss the results on individual and household characteristics – income, education, race, and household structure in on financial market participation. These results are presented in Table 3. In general, the effect of individual and household level variables on savings account and checking account ownership are similar.

Income has a strong positive effect on financial market participation. If monthly per capita household income were to increase by one standard deviation from its mean, by \$2,764, the likelihood of savings account ownership would increase by 12 percentage points, a 23% increase relative to the observed percentage of the individuals in the sample who have a savings account of 53%. Similarly, checking account ownership would increase by 12 percentage points, and this represents a 30% increase relative to the observed likelihood of owning a checking account of 35%. Being unemployed or out of the labor force has a strong negative impact on savings account ownership, but a small positive impact on the probability of owning a checking account. The different effect of age and labor market status on savings and checking account ownership is most likely driven by greater ownership of interest-bearing checking accounts among retirees.

Educational attainment plays a very important role in explaining patterns of financial market participation. For example, compared to those with less than a high school diploma, high school graduates are about 13 percentage points more likely to own a savings account and 17 percentage points more likely to have a checking account. Individuals who have completed some college are 21 percentage points more likely to have a savings account and 26 percentage points more likely to have a checking account compared to those who did not complete high school. The predicted gap in account ownership between college graduates and those who did not complete high school is even larger, 24 percentage points for savings accounts and 35

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<sup>9</sup> While household wealth may provide a more suitable measure of permanent income or the lifetime resources for a given household, the SIPP wealth variable is only available in the topical modules (and is measured every 8 months).

<sup>10</sup> We should note that we find similar results on stock, individual retirement account, and mutual fund ownership (and these results are shown in Appendix I). We omit a detailed discussion of these results in the interest of space.

percentage points for checking accounts. The figures are similar when we compare individuals with an advanced degree to individuals who did not complete high school.

Older individuals are more likely to own checking accounts. There are some non-linearities with respect to the effect of age on savings account ownership. While age is negatively associated with savings account ownership, age squared has a positive and significant impact on savings account ownership. Being married has a large positive impact on savings and checking account ownership, increasing the probability of savings account ownership by more than 20 percentage points and the likelihood of checking account ownership by 17 percentage points. Interestingly, men are significantly less likely to report holding a savings and checking accounts. We also note that, compared to whites, non-whites are 11 – 12 percentage points less likely to have a savings or a checking account. The number of children in the household reduces the likelihood of having a savings or a checking account by about 2 percentage points for each additional child.

### *Impact of Immigrant Status*

We now turn to the key variable of interest. Immigrants are significantly less likely to participate in financial markets, compared to the native-born. Specifically, immigrants are 7.4 percentage points less likely to own a savings account compared to a similar native-born individual. Immigrants are also 6.1 percentage points less likely to own a checking account compared to a similar native-born individual. Across a range of additional financial assets— notably, stock, mutual funds, and individual retirement accounts, we note that immigrants have a lower likelihood of financial asset ownership, compared to the native born. From Appendix I, immigrants are 5.6 and 4 percentage points less likely to own stock and individual retirement accounts, respectively compared to a similar native-born individual.

In Table 4, we consider the role of time in the U.S. on the financial market participation of immigrants relative to the native-born. Specifically, we estimate the additional effect of being a recent immigrant on savings account ownership (column 1) and checking account ownership (column 3). We define recent immigrants to be those who arrived in the U.S. in 1990 or more recently. At most they would have lived in the U.S. for six years at the beginning of the SIPP survey. Columns (2) and (4) include a full-set of year of arrival controls and allow us to consider

how the impact of being an immigrant on savings and checking account ownership, respectively, varies more generally with time in the U.S.<sup>11</sup>

While immigrants as a group are 7.4 percentage points less likely to have a savings account and 6 percentage points less likely to have a checking account compared to the native-born, recent immigrants are 18 percentage points less likely to have a savings account and 12 percentage points less likely to have a checking account (see Table 4, columns 1 and 3).<sup>12</sup> Recent immigrants are particularly likely to differ in important ways from the native-born in their familiarity and knowledge of U.S. financial markets. English language ability and legal status are likely to be important concerns for recent immigrants compared to their more established counterparts. In addition, information costs may impose significant barriers to immigrant participation in formal financial markets although these costs would tend to decrease as immigrants gain U.S. experience. The estimates presented in columns 2 and 4 suggest that this is indeed the case. While immigrants who arrived between 1990 and 1996 are 18 percentage points less likely to have a savings account and 12 percentage points less likely to have a checking account, immigrants who arrived between 1985 and 1989 are only 9 percentage points less likely to have a savings account and only 8 percentage points less likely to have a checking account, compared to the native-born.

With one exception, the cohort controls are not significantly different from zero for immigrants who arrived before 1985, suggesting that partial financial market assimilation may happen in the first ten to fifteen years after migration and then stops. Interestingly, we find that immigrants who arrived between 1975 and 1979 are as likely as the native born to have a savings account.<sup>13</sup> Altogether, the estimates presented in Table 4 indicate that immigrant financial market assimilation is partial at best. Taking into account U.S. experience and a rich set of controls, immigrants are about 5 percentage points less likely to have a savings account or a checking account compared to the native-born.

## **B. Decomposing the Immigrant-Native Gap in Financial Market Participation**

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<sup>11</sup> In addition to controls for being an immigrant and duration of stay in the U.S., the estimates presented in Table 4 also contain the same set of control variables that were included in the estimates presented in Table 3.

<sup>12</sup> For stock and IRA ownership, recent arrivals are 12 percentage points and 17 percentage points less likely to own these assets compared to similar native-born.

<sup>13</sup> The 1975-1979 cohort may have been particularly impacted by the 1986 Immigration Reform and Control Act which provided amnesty in the form of legal permanent residence for undocumented immigrants who could prove that they had been living continuously in the U.S. prior to January 1, 1982. Agricultural workers who had worked in the U.S. for at least 90 days in the year prior to May 1, 1986 were also eligible for amnesty.

Having documented that there is an important gap in immigrant-native financial market participation, we turn now to quantifying the fraction of the gap that can be explained by characteristics and by returns to characteristics (or “prices”). Specifically, we are interested in quantifying the key characteristics that drive the portion of the gap that can be attributed to group differences in characteristics, and quantifying the relative importance of group differences in education, income, and metropolitan areas in explaining immigrant-native gaps in participation in financial markets. Given the non-linearity of the logit equation, we use a variation of the Blinder-Oaxaca decomposition (Blinder, 1973; Oaxaca, 1973), which is described in Fairlie (2003).

Table 5 summarizes the nonlinear decomposition of the immigrant-native gap in financial market participation based on Fairlie (2003). The decomposition results presented in Table 5 suggest that group differences in characteristics and the returns to characteristics between immigrants and the native-born are equally important in explaining the gap in financial market participation for the two groups.<sup>14</sup> We also consider the role that specific characteristics play in creating the observed differences between immigrants and the native-born. As one might expect, education and income differences between immigrants and the native-born play a key role in increasing the gap in financial market participation that is due to characteristics. According to our decomposition results, individual, family, and MSA characteristics account for about 50 to 70% of the difference in financial market participation that can be attributed to characteristics. Interestingly, differences in the metropolitan areas where immigrants and the native-born live play an important role, accounting for about 17 percentage points of the gap that is due to characteristics. This suggests that on average, the financial market participation of immigrants would be higher if they lived in the same MSAs as the native-born.

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<sup>14</sup> For the logit equation, the decomposition of the native/immigrant gap is expressed below.  $F(\cdot)$  is the cumulative distribution function from the logistic distribution,  $X^m$  is a row vector of average values for the individual characteristics and MSA effects,  $\hat{\beta}^m$  is a vector of coefficient estimates for group  $m$ , and  $Y^m$  is the average probability of owning an account for group  $m$ . We present the decomposition using immigrant coefficients in the first term:

$$\bar{Y}^N - \bar{Y}^I = \left[ \sum_{i=1}^{N^I} \frac{F(X_i^I \hat{\beta}^I)}{N^I} - \sum_{i=1}^{N^I} \frac{F(X_i^I \hat{\beta}^I)}{N^I} \right] + \left[ \sum_{i=1}^{N^N} \frac{F(X_i^N \hat{\beta}^N)}{N^N} - \sum_{i=1}^{N^N} \frac{F(X_i^N \hat{\beta}^I)}{N^N} \right]$$

### **C. Unobserved Heterogeneity and Financial Market Participation**

The decomposition results presented above are useful in quantifying the separate role that group differences in measurable characteristics play in explaining differences in the financial behavior of immigrants and the native-born. However, while observed characteristics such as household income, race and ethnicity, education, age, and family structure play an important role, it is also likely that unobservables including tastes and preferences and the degree of risk aversion, may differ across otherwise similar immigrants and the native-born. While we cannot observe all these characteristics directly, various empirical techniques help us examine the extent to which differences in immigrant-native financial market behavior are driven by unobserved heterogeneity.<sup>15</sup>

We take two approaches to dealing with unobserved heterogeneity. First we investigate the impact of adding control variables to the estimates presented in Table 3 in an effort to better account for omitted variables. We are concerned that to the extent that omitted variables are correlated with being an immigrant, they will bias the coefficient on the immigrant indicator variable in the baseline estimates of financial market participation. In the estimates presented in Table 6 and discussed in sub-section [1] below, we explore the role of ethnicity, legal status, language, and other potential sources of bias. In addition, we make use of the panel nature of the SIPP data and estimate transitions into and out of financial market ownership. The estimates of changes in financial market behavior from one period to the next, account for unobserved heterogeneity by implicitly differencing out the effect of fixed individual characteristics. If being an immigrant has a similar effect on owning a savings or a checking account as it does on transitions in ownership, then we gain confidence that our baseline findings are not overly influenced by unobserved heterogeneity. These estimates are presented in Tables 7A (Entry) and 7B (Exit) and discussed in sub-section [2] below.

#### **1. Unobserved Heterogeneity – Additional Control Variables**

Before discussing the results which include additional controls, it is useful to reconsider the estimates which include year of arrival controls in the light of unobserved heterogeneity

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<sup>15</sup> In particular, financial support of relatives in the country of origin and the use and availability of informal substitutes for formal financial products and services are also likely to be important sources of unobserved heterogeneity.



(Table 4). To some extent, potential biases in the effect of being an immigrant on financial market participation due to unobserved heterogeneity are addressed in these estimates. By including year of arrival controls, the influence of omitted variables (such as English language ability and legal status which tend to vary over time and by year of arrival cohort) will tend to show up in the coefficients on the year of arrival controls and will reduce the coefficient on immigrant status. We find that including the year of arrival controls reduces the impact of being an immigrant on financial market participation from negative 7 percent to negative 4 percent for savings and from negative 6 percent to negative 4 percent for checking.

In Table 6, we investigate the effect of specific omitted variables on the financial market participation of immigrants relative to the native-born. While we are interested in the direct effect of the additional control variables, we are also interested in how much the coefficient on immigrant status changes as a result of adding controls.

For comparison purposes, the baseline results from Table 3 are presented in column (1) of Table 6. The first source of unobserved heterogeneity that we consider is racial differences within the immigrant community. This estimate addresses the possibility that non-white immigrants differ significantly in their use of (or, potentially, access to) formal financial institutions compared to that of white immigrants because of discrimination by financial institutions or beliefs about discrimination by financial institutions. Recent empirical studies of household financial behavior have documented significant differences in the use of financial services by race, even after controlling for income and education (Altonji and Doraszelski, 2005; Blau and Graham, 1990; Chiteji and Stafford, 1999). In column (2) we allow the effect of race to differ for immigrants and the native-born. In the baseline estimates, the effect of being “non-white” is restricted to be the same for immigrants and the native-born. We find relatively small, but significant differences in the financial market behavior of immigrants by race. We should note that adding the interaction of immigrant status and race actually increases the negative effect of being an immigrant on financial market participation.

In column (3), we consider the effect of legal status at the time of migration on financial market participation. Immigrants who lack the legal right to live and work in the U.S. may face additional barriers to opening a savings or checking account. Many financial institutions, particularly during the survey period, required a social security number and a U.S. Driver’s

License to open an account.<sup>16</sup> While the SIPP data do not include information on whether an immigrant is undocumented upon arrival or at the time of the survey, they do report whether an immigrant was a legal permanent resident at the time of migration. Our results suggest that permanent residence has a positive and significant impact for both savings and checking account ownership. Immigrants who arrived in the U.S. as permanent residents are about 2 to 3 percentage points more likely to own savings and checking accounts, compared to other immigrants. However, adding the legal status variable does not significantly reduce the negative effect of being an immigrant on financial market participation.

The baseline estimates of financial market participation include controls for education and assume that education has the same impact on financial market participation for immigrants and the native-born. In column (4) of Table 6 we consider the possibility that the impact of being an immigrant on financial market participation varies with education *among* the immigrant population. If immigrants with exposure to higher education (beyond high school) also have better employment opportunities, enhanced English skills, and access to different sources of information about financial markets, their behavior may differ significantly from less-educated immigrants. We allow for this possibility by adding an interaction term to the set of control variables: immigrant multiplied by a variable that is equal to one if an individual has completed more than a high school education.

We find that immigrants with more than a high school degree are 16 percentage points more likely to have a savings account and 25 percentage points more likely to have a checking account compared to immigrants who have at most completed a high school degree. Among the native-born, those who have a high school degree or more are 13 percentage points more likely to have a savings account and 15 percentage points more likely to have a checking account. Education appears to have a bigger impact on immigrant financial market behavior than it does on native financial market behavior. This suggests that education captures additional aspects of the immigrant experience like access to job sources, English language ability, and information about financial products and services. For immigrants as a whole, however, adding the interaction of immigrant with a high school education or greater makes the contrast between immigrant and native financial market participation even starker: immigrants are 10 percentage

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<sup>16</sup> Although many U.S. financial institutions require a Social Security number in order to open an account, a growing number of banks accept Individual Taxpayer Identification Numbers (ITINs) as an alternative and recognize identification cards issued by consular offices of the immigrant's country of origin.

points less likely to have a savings account and 11 percentage points less likely to have a checking account in these estimates. It appears that failing to allow education to have a different effect immigrants and the native-born in the baseline estimates led to a downward bias in the estimated impact of being an immigrant on financial market participation.

In column (5), we repeat the estimation of the baseline specification on a sample that excludes Mexican immigrants. Mexican immigrants make up approximately one-third of the immigrant sample and have some distinguishing characteristics that are difficult to measure in the SIPP data and that are also potential sources of bias. Specifically, Mexican immigrants are more likely to be undocumented. They also have higher propensities for return migration compared to other immigrants and this may have implications for savings behavior (see Dustmann, 1997; Galor and Stark, 1990), for example). In addition, Mexican immigrants tend to have lower English ability and education compared to other immigrants. Eliminating this immigrant group from the sample does not substantively alter the conclusions of the baseline estimates. Excluding the Mexican sample, we find that immigrants are 5 percentage points less likely than natives to have a savings account (compared to 7 percentage points in the baseline case) and 5 percentage points less likely to have a checking account (compared to 6 percentage points in the baseline estimates).

In column 6, we restrict the sample to native and immigrant Hispanics. Consistent with several studies that have documented low rates of financial asset holdings among Hispanic immigrants compared to the Hispanic native-born, we also find low rates of savings account ownership among Hispanics compared to other ethnic groups (Cobb-Clark and Hildebrand, 2006c; Smith, 1995). In particular, we find that ownership rate of savings accounts for Hispanic immigrants is low at 19% when compared to ownership rates of 38% for native-born Hispanics. For checking accounts, Hispanic immigrants ownership rates are 10% compared to much higher ownership rates of 28% among Hispanic native-born individuals. However, for Hispanics, taken as a group, it is not clear how much of their lower participation rates can be explained by immigrant status and how much can be explained by English language proficiency and other barriers. While information on English language proficiency is not available in the SIPP data, we can learn about the relative importance of language proficiency (compared to other factors) by restricting our sample to Hispanics. When we restrict our sample to Hispanics, we still find significant differences, of roughly the same magnitude as the baseline estimates, in financial

market participation between native-born Hispanics and Hispanic immigrants. Hispanic immigrants are 6 percentage points less likely to have a savings account and a checking account compared to native-born Hispanics. These estimates increase our confidence that the baseline estimates of the gap in immigrant-native financial market behavior is not driven by omitted variables like English language ability.

We have examined a number of potential sources of bias in the baseline results and found that they are robust to adding additional controls for race, legal status, and education and also to studying a sample which exclude Mexican immigrants and a sample made up solely of Hispanic immigrants and natives. If anything, adding controls for race, legal status, and education widens the gap in the predicted financial market participation of immigrants and natives. The estimates of the gap in financial market behavior derived from the sample which excludes Mexicans and from the sample of all Hispanics are similar in magnitude and substance to the baseline results. Unobserved heterogeneity along the dimensions discussed above does not seem to account for the gap in immigrant-native financial market participation.

## **2. Unobserved Heterogeneity – Entry into and Exit out of Account Ownership**

In Tables 7A and 7B, we estimate transitions into and out of account ownership. These estimates are of interest for at least two reasons. First, they offer additional insights into why immigrant financial behavior differs from that of natives. If differences in behavior are driven by differences in the propensity to enter into account ownership, then one reason for immigrant-native differences may lie in differential access to information about financial services and products that impacts the decision to open an account. If the gap is driven by differences in the likelihood of closing an account, then lower financial market participation among immigrants may be driven by increased vulnerability to economic shocks and the presence of informal financial services. A second motivation for examining transitions into and out of account ownership is that these estimates provide another means for controlling for unobserved heterogeneity. Since the dependent variable in these estimates reflects changes in financial market decisions, the impact of time-invariant individual characteristics (tastes and preferences, in particular, risk aversion, unobserved ability, home country experiences, English language proficiency, propensity for return migration or for private transfers to relatives living outside the U.S., for example) have been implicitly differenced out.

For estimates of entry, the dependent variable is equal to one if the individual reports owning an account at time  $t$  and not owning an account at time  $t - 1$ .<sup>17</sup> The dependent variable is equal to zero if the individual reports no account ownership at time  $t$  and at time  $t - 1$ . All of the transition estimates include the explanatory variables described in the discussion of the baseline results in Table 3. Our estimates are based on logit maximum likelihood specification and standard errors are adjusted to allow for correlation across observations at the individual level.

Estimates of transitions **into** account ownership are found in panels A and B of Table 7 for savings and checking accounts, respectively. There are two estimates of entry into account ownership. The estimates presented in column (1) of Table 7A include a control for being an immigrant. In column (2) an additional control for being a recent immigrant is added. The estimates also include all of the control variables described above in the discussion of Table 3.

From column (1) we see that immigrants are less likely to enter into savings and checking account ownership, although the effect is not statistically significant. The likelihood of opening a checking account is predicted to be lower for immigrants, but only significant at the 11% level of significance. The estimates presented in column (2) suggest that the differences in the likelihood of opening savings and checking accounts for immigrants and natives is driven by recent immigrants, who are 1 percentage point less likely to open a savings account and 0.7 percentage points less likely to open a checking account.

In panels C and D of Table 7 we present estimates of the likelihood of transitions **out of** account ownership (**exits**) for savings accounts (panel C) and checking accounts (panel D). The dependent variables in panels C and D are equal to one if an individual, who had an account at time  $t - 1$ , reports not having an account at time  $t$ . The dependent variable (exit) is equal to zero if the individual reports having an account at both  $t - 1$  and  $t$ . Individuals with no account at both  $t-1$  and  $t$  are excluded from the analysis.

From column (1) we see that immigrants are 1.4 percentage points more likely to exit from both savings and checking account ownership. This corresponds to a 27% higher likelihood of closing a savings account or a checking account for immigrants compared to natives, relative to the observed frequency of savings account closures of 5.2% and the observed

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<sup>17</sup> We should note that the entry estimates are restricted to those who report no ownership at time  $t - 1$  and that the exit estimates are restricted to those who do own an account at time  $t - 1$ .

frequency of checking account closures of 5.3%. In contrast to the estimates for opening accounts, the difference in account closures for immigrants and natives is *not* driven by recent immigrants. The estimates in column (2) show that being a recent immigrant has no significant additional impact on the likelihood of closing an account compared. The coefficient on being an immigrant remains basically unchanged in size and significance when the recent immigrant control is included.

Concerns about bias due to unobserved heterogeneity in the baseline estimates are mitigated by the fact that we see a similar effect of being an immigrant on estimates of transitions into and out of account ownership. In addition, the transition estimates suggest that the underlying causes of differences in financial market participation among immigrants and natives are likely to differ for recent and established immigrants. Recently arrived immigrants are less likely to open accounts than both natives and more established immigrants, which is consistent with barriers to information and the likelihood of return migration limiting entry into mainstream financial markets. Information barriers and the potential for return migration do not seem to limit the entry of more established immigrants, however. In contrast, all immigrants, regardless of how recently they arrived in the U.S., are more likely to close savings and checking accounts compared to the native-born. It seems highly unlikely that this effect is driven by information, since this group had enough information about U.S. financial services to open an account in the first place. The fact that immigrants are more likely to close accounts suggests that part of the explanation for differences in the financial behavior of immigrants and natives may lie in their relative vulnerability to economic shocks. One possibility here is that adverse economic circumstances force immigrants to liquidate accounts more frequently than they do natives. This could be due to the fact that immigrants are over-represented in sectors of the economy – agriculture and services, for example – that are particularly cyclical. Another potential explanation could be that immigrants rely on informal financial markets or are more likely to be subject to adverse shocks compared to the native born because they provide economic support to more people, including family members who live in their country of origin. In contrast, the family members of the native-born are more likely to be covered by U.S. social safety net programs.

#### **D. Location and Financial Market Participation**

A final issue in our analysis is how location contributes to immigrant-native differences in financial market behavior. There are several reasons for focusing on location (MSA of residence) in seeking to better understand why immigrants make different financial decisions compared to otherwise similar native-born individuals. First, the decomposition exercise presented in Table 5 found that if immigrants lived in the same metropolitan areas (MSAs) as the native-born, the explained difference in immigrant-native financial market participation would fall by about 17%. Second, many other researchers have found important effects of residential settlement on immigrant behavior. For example, Borjas (1998, 2000) finds that immigrants who live in ethnic enclaves have lower wage growth and greater income uncertainty. The geographic clustering of immigrants has also been shown to lower educational attainment and language proficiency (see Gang and Zimmerman, 2000; Chiswick and Miller, 2002).<sup>18</sup>

We examine the possibility that an immigrant who lives in an urban community where there is a high concentration of people who have emigrated from the same country may differ in financial market behavior from an immigrant from the same country who lives in a community with a lower concentration of people who emigrated from the same country. Focusing on a measure of the network that immigrants are likely to interact with seems very reasonable given our interest financial behavior as a number of recent studies have shown that social interactions have important effects on financial decisions.<sup>19</sup> One testable hypothesis here is that low financial market participation of immigrants may be reinforced when immigrants have a large network of individuals to interact with who came from the same country of origin. Related to this, ethnic networks may also provide information about and be sources of informal alternatives for formal financial services.

To measure ethnic concentration, we supplement the SIPP data with information from the 1990 Integrated Public Use Microdata Sample as (IPUMS) 1% sample of the U.S. Census to construct the fraction of a given MSA population that was born in a specific country. Summary information about the Ethnic Concentration variable is found in Appendix II. For an immigrant from a given country,  $k$ , (for example, Mexico), ethnic concentration is measured as the share of

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<sup>18</sup> Munshi (2003) shows that ethnic networks affect employment opportunities for Mexican immigrants even after instruments for network size are used.

<sup>19</sup> For example, Hong, Stein, and Kubik (2004) show that social interactions have important effects on stock market participation. Similarly, Madrian and Shea (2000) and Duflo and Saez (2003) show that decisions to participate in employer-sponsored retirement plans are influenced by the choices of co-workers.

immigrants from that country (for example, Mexico) that reside within a given MSA,  $j$ . More formally, we define ethnic concentration for country  $k$  and MSA  $j$  as follows:

$$Ethnic\ Concentration_{kj} = \frac{\# \text{ of individuals born in country } k \text{ residing in destination community } j}{\text{Total } \# \text{ of individuals (including natives) residing in destination community } j}$$

Estimates which include the ethnic concentration variable are found in Table 8. Column (1) adds this variable to the baseline estimates for savings (in panel A) and checking (in panel B). In column (2) we also add a control for immigrants who have arrived in the U.S. since 1990. In column (3), the interaction of the arrival variable and the ethnic concentration is also included in the estimation. As always, the estimates in columns (1) – (3) also include MSA fixed-effects.

The results in Table 8 provide evidence that patterns of residential settlement may play an important role in understanding immigrant participation in U.S. financial markets. We find that the size of the ethnic network has a significant negative impact on financial market participation. Immigrants who live in MSAs with higher ethnic concentrations are less likely to use mainstream financial services. In order to quantify these effects, we consider the case of Mexican immigrants living in the Chicago and Milwaukee MSAs. The Milwaukee MSA is located only 90 miles north of Chicago. However, Mexican immigrants account for 4.22% of the population in the Chicago MSA (which is the highest representation among immigrants in Chicago) while Mexican immigrants have the second highest representation among immigrants in Milwaukee and account for 0.51% of the population. Across all of the MSAs in the sample, the average concentration of Mexicans is 2.61%, and this ranges from a low of 0.01% to a high of 33.04%.

Looking at column (1) of Table 8, we see that immigrants in general are 6 percentage points less likely than the native-born to have a savings account. For Mexican immigrants living in Milwaukee, the community characteristics do not change this figure too much: they are an additional 0.31 percentage points less likely to have savings account. However, Mexican immigrants living in Chicago are an additional 2.5 percentage points less likely to have a savings account. Holding other variables fixed, if a Mexican immigrant moved from Chicago to Milwaukee, the likelihood that they would have a savings account would go up by 2.2 percentage points.<sup>20</sup>

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<sup>20</sup> For checking account ownership, the overall effect of being an immigrant is somewhat smaller. Immigrants overall are predicted to be 3 percentage points less likely to have a checking account. However, the magnitude of



In column (2) of Table 8, we add an additional control variable for being a recent immigrant. As we have seen before, recent immigrants are much less likely than similar natives to own a savings account or a checking account. In these estimates, recent immigrants are predicted to be 16 percentage points less likely to have a savings account and 10 percentage points less likely to have a checking account. Adding the recent immigrant control variable does not appreciably change the size or the significance of the ethnic concentration variable, however.

In column (3) of Table 8, we consider the possibility that the impact of living among a substantial population of immigrants from the same origin country may differ for recent and more established immigrants. We find evidence that this is in fact the case. The size of the potential immigrant network appears to have an important effect on the financial market participation of immigrants, particularly for recent immigrants who may be especially reliant on other immigrants who share the same country of origin for information about U.S. financial markets. This is consistent with the finding in Table 7 that recent immigrants are less likely to open savings and checking accounts compared to the native-born, but that more established immigrants behave similarly to natives when it comes to opening accounts.<sup>21</sup>

Compared to our baseline findings, the effect of being an immigrant is lower in the estimates that include the ethnic concentration variable. For savings account ownership, the effect of being an immigrant is 56 – 81 percent lower according to the estimates that include the ethnic concentration variable compared to the analogous estimates which do not control for ethnic concentration. For checking account ownership, the impact of being an immigrant is estimated to be 39 – 49 percent lower when the ethnic concentration variable is included. According to these estimates, somewhere between 20 and 60 percent of the effect of being an immigrant may operate through residential settlement.

By including MSA fixed-effects in the estimates, we address the concern that residents of a given community share a common economic environment, or have similar preferences. For example, there may be a lower supply of financial services or limited employment prospects in

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ethnic concentration variable is larger. Mexican immigrants living in Milwaukee are an additional 5.2 percentage points more likely to have a checking account compared to Mexican immigrants living in Chicago.

<sup>21</sup> Estimates of financial market transitions that include ethnic concentration (available from the authors) reinforce the message that the entry behavior of recent immigrants, but not more established immigrants, is influenced by patterns of residential settlement. Exit from savings account ownership is higher for immigrants but does not vary with ethnic concentration. In contrast, exit rates from checking account ownership are higher for immigrants and go up with ethnic concentration. We note that these estimates come closest to dealing with unobserved heterogeneity, including decisions about where to live and country of origin effects, because they implicitly difference out fixed characteristics.

one MSA compared to another. MSA fixed effects do not, however, capture all forms of unobserved heterogeneity such as variation in the supply of financial services or employment prospects by country of origin within a MSA.<sup>22</sup> Moreover, estimates of the impact of residential settlement on financial market participation may be biased because the decision about where to live is unlikely to be random.<sup>23</sup> For example, it is quite possible that immigrants who choose to live in Milwaukee differ in some unmeasured way from those who choose to live in Chicago and that the characteristics that impact the choice about where to live also impact financial market behavior.

These estimates do not tell us the exact mechanism through which ethnic concentration impacts financial market participation. It is certainly possible that there is a direct effect of ethnic concentration on financial market participation. Specifically, immigrants residing in ethnically concentrated locations are more likely to interact with and get information about financial products and services from other immigrants from the country of origin and this reinforces already low levels of formal financial market participation among this group. It is also possible that there is an indirect effect of ethnic concentration on financial market participation. For example, as noted above, other researchers have found that ethnic concentration reduces immigrant language acquisition, raises income uncertainty, lowers wage growth, and reduces human capital accumulation. While we are able to hold education and income constant in our estimates, we do not have data on language proficiency or income uncertainty, so the coefficient on the ethnic concentration variable will capture both direct and indirect effects. If there is an indirect effect it would mainly operate through language or income uncertainty, since the estimates control for education, income, and employment status.

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<sup>22</sup> If country of origin characteristics also influence the choice of destination community (as in Bauer et al, 2005) then the unobserved determinants of immigrant location choice are likely to vary by country of origin within a given MSA. However, we find that effect of ethnic concentration on financial market behavior is robust when we interact country characteristics (such as the level of financial development in the country of origin) with ethnic concentration in a given MSA. This specification allows us to include both MSA and country fixed effects.

<sup>23</sup> Bauer, Epstein, and Gang (2005) find that as immigrants gain English language proficiency they choose communities with smaller ethnic networks, and Bartel (1989) finds that skilled immigrants are less geographically concentrated than their unskilled counterparts.

## V. Conclusions

This paper seeks to add to our existing knowledge on the prospects for immigrant wealth assimilation, and immigrant assimilation more generally, by studying the financial market behavior of U.S. immigrants and comparing it to the native-born. Compared to similar natives, immigrants are less likely to participate in financial markets. We show that lower rates of financial market participation tend to persist even for immigrants who have lived in the U.S. for several years, compared to the native-born. In addition, immigrant status has a significant impact on transitions into and out of account ownership. Specifically, immigrants are somewhat less likely to open accounts and more likely to close accounts compared to similar native-born individuals. Concerns that the results are driven by unobserved heterogeneity are reduced, because the effect of being an immigrant is similar for financial market participation and for changes in financial market participation.

Our results suggest that a large share of the immigrant-native gap in financial market participation is driven by education, income, and geographic location. We present some evidence that the explanation for differential behavior of immigrants relative to natives has to do with variations in patterns of residential settlement, specifically ethnic concentration within a given MSA. Our results on entry into account ownership are consistent with social interaction effects, in which immigrants, particularly recent arrivals, have fewer connections with mainstream society and lack information about formal financial markets. However, the finding that immigrants have higher exit rates from account ownership suggests that the informational hypothesis cannot be the sole explanation for low rates of immigrant participation in mainstream financial markets. Past research has shown that immigrant residing in ethnically concentrated areas have low levels of English proficiency and higher income uncertainty. Thus an additional channel through which ethnic concentration may affect financial participation is through greater labor market insecurity and greater language barriers among immigrants residing in ethnic enclaves. Finally, immigrants residing in ethnically concentrated areas may have access to informal alternatives to formal financial services.

Our findings on ethnic concentration are intriguing in light of a growing number of studies that have shown that social interactions play an important role in many economic decisions, including financial market participation, welfare usage, and criminal behavior. An important goal of future research in this area is to identify the precise mechanism through with

ethnic concentration affects immigrant behavior – controlling in particular for factors that may influence financial market indirectly through location choice. Understanding the mechanism through which ethnic concentration impacts immigrant behavior may have important policy implications. For example, if ethnic concentration mainly affects financial market participation through word-of-mouth learning about mainstream financial services, then financial literacy programs may have large multiplier effects within immigrant populations. However, to the extent that ethnic enclaves provide immigrants with informal alternatives to formal financial markets, then additional research may be needed to understand the factors that increase the attractiveness informal financial services for immigrants, compared to similar natives. Because financial transactions rely on trust and confidence in institutions, the financial market behavior of immigrants can provide key insights into the process of immigrant adaptation to U.S. social and economic life.

## REFERENCES

- Altonji, Joseph A. and Ulrich Doraszelski. 2005. "The Role of Permanent Income and Demographics in Black/White Differences in Wealth," *Journal of Human Resources* 40(1):1-30
- Amuedo-Dorantes, Catalina, Cynthia Bansak and Susan Pozo. 2005. "On the Remitting Patterns of Immigrants: Evidence from Mexican Survey Data," *Economic Review, Federal Reserve bank of Atlanta*, 1<sup>st</sup> Quarter.
- Amuedo-Dorantes, Catalina and Susan Pozo. 2002. "Precautionary Saving by Young Immigrants and Young Natives," *Southern Economic Journal*, 69(1):48-71.
- Bartel, Ann P. 1989. "Where Do the New Immigrants Live?," *Journal of Labor Economics*, 7(4):371-391.
- Bauer, Thomas, Gil S. Epstein, and Ira N. Gang, 2005. "Enclaves, Language and the Location Choice of Migrants," *Journal of Population Economics*, 18(4): 649-662
- Bender, Keith A., 2004. "The Well-Being of Retirees: Evidence Using Subjective Data" *Boston College Center for Retirement Research Working Paper #2004-24*
- Blau, Francine D. and John Graham, 1990. "Black-White Differences in Wealth and Asset Composition," *Quarterly Journal of Economics*, 105(2):321-39.
- Blinder, Alan S. 1973. "Wage Discrimination: Reduced Form and Structural Estimates," *Journal of Human Resources*, 8(4):436-455.
- Borjas, George 2002. "Homeownership in the Immigrant Population," *Journal of Urban Economics* 52(3): 448-476
- Borjas, George J. 2000. "Ethnic Enclaves and Assimilation", *Swedish Economic Policy Review*: 89-122
- Borjas, George J., 1998. "To Ghetto or Not to Ghetto? Ethnicity and Residential Segregation," *Journal of Urban Economics* 44: 228-253.
- Borjas, George J. 1995 "Ethnicity, Neighborhoods, and Human-Capital Externalities", *American Economic Review*, 85(3): 365-90.
- Borjas, George J. 1987. "Self-Selection and the Earnings of Immigrants," *American Economic Review*, 77(4):531-53.
- Borjas, George J. 1985. "Assimilation, Changes in Cohort Quality, and the Earnings of Immigrants," *Journal of Labor Economics*, 3(4): 463-489.
- Carroll, Christopher D., Rhee, Byung-Kun and Changyong Rhee, 1999. "Does Cultural Origin Affect Saving Behavior? Evidence from Immigrants", *Economic Development and Cultural Change*, 48(1): 33-50.
- Chiswick, Barry R. and Paul W. Miller 2002 "Immigrant Earnings: Language Skills, Linguistic Concentrations and the Business Cycle", *Journal of Population Economics*, 15(1): 31-57

- Chiswick, Barry R and Paul W. Miller. 1996 "Ethnic Networks and Language Proficiency among Immigrants in the United States," *Journal of Population Economics*, 9(1): 19-35
- Chiswick Barry R. 1978 "The Effect of Americanization on the Earnings of Foreign-born Men" *Journal of Political Economy*, 86(5): 897-921.
- Chiteji, Ngina S. and Frank. P. Stafford. 1999. "Portfolio Choices of Parents and Their Children as Young Adults: Asset Accumulation by African-American Families," *American Economic Review*, 89(2): 377-80
- Cobb-Clark, Deborah and Vincent Hildenbrand, 2006a. "The Wealth and Asset Holdings of U.S. and Foreign Born Households: Evidence from SIPP data," *Review of Income and Wealth* 52(1)17-42
- Cobb-Clark, Deborah and Vincent Hildenbrand, 2006b. "The Wealth of Mexican Americans," *Journal of Human Resources* forthcoming
- Cobb-Clark, Deborah and Vincent Hildenbrand, 2006c. "Social Science Quarterly," forthcoming
- Cobb-Clark, Deborah and Sherrie Kossoudji 2002. "Coming Out of the Shadows: Learning about Legal Status and Wages from the Legalized Population." *Journal of Labor Economics*, 20(3):598-628
- Duflo, Esther and Emmanuel Saez 2003 "The Role of Information and Social Interactions in Retirement Plan Decisions: Evidence from a Randomized Experiment", *Quarterly Journal of Economics*, 118(3): 815-42
- Duncan, Greg J. and Jeanne Brooks-Gunn, 1997 (eds). Consequences of Growing Up Poor, New York: Russell Sage.
- Dustmann, Christian, 1997. "Return Migration, Savings, and Uncertainty," *Journal of Development Economics*, 52: 295-316
- Fairlie, Robert W. 2003. "An Extension of the Blinder-Oaxacca Decomposition Technique to Logit and Probit Models," *Economic Growth Center Yale University*, 17(1), Center Discussion Paper No. 873.
- Fairlie, Robert W. 1999. "The Absence of the African-American Owned Business: An Analysis of the Dynamics of Self-Employment," *Journal of Labor Economics*, 17(1), 80-108.
- Gang, Ira N. and Zimmermann, Klaus F. 2000. "Is Child like Parent? Educational Attainment and Ethnic Origin", *Journal of Human Resources*, 35(3): 550-69
- Galor, Oded and Stark, Oded 1990. "Migrants' Savings, the Probability of Return Migration and Migrants' Performance" *International Economic Review*, 31(2): 463-67.
- Guiso, Luigi, Sapienza, Paola, and Zingales, Luigi. 2004. "The Role of Social Capital in Financial Development," *American Economic Review*, 94 (3):526-556.
- Hao, Lingxin 2004 "Wealth of Immigrant and Native-Born Americans." *International Migration Review* 38(2): 518-546
- Hao, Lingxin 2001 "Immigrants and Wealth Stratification in the U.S." *Joint Center for Poverty Research Working Paper* #216

- Haurin, Donald R., Patric Hendershott, and Susan Wachter, 1997. "Borrowing Constraints and the Tenure Choice of Young Households" *Journal of Housing Research* 9(1):137-154.
- Holtz-Eakin David, Joulfaian, David and Harvey S. Rosen, 1994. "Sticking it Out: Entrepreneurial Survival and Liquidity Constraints," *The Journal of Political Economy*, 102(1): 53-75
- Hong, Harrison, Stein, Jeremy, and Jeffrey Kubik 2004. "Social Interaction and Stock-Market Participation," *The Journal of Finance* 59(1):137-163
- Krivo, Laurie and Robert L. Kauffman, 2004. "Housing and Wealth Inequality: Racial-ethnic Differences in Home Equity in the United States," *Demography* 41(3):585-605.
- Light, Ivan 1972. Ethnic Enterprise in America Berkeley: University of California Press.
- Madrian, Brigitte and D. Shea, 2000. "Peer Effects and Savings Behavior in Employer-Sponsored Savings Plans," *University of Chicago Working Paper*
- Mayer, Susan E. 1997. What Money Can't Buy: Family Income and Children's Life Chances Cambridge, MA: Harvard University Press.
- Modigliani, Franco and Richard Brumberg 1954. "Utility Analysis and the Consumption Function: An Interpretation of Cross-Section Data," In: edited by Kurihara, K.K: Post-Keynesian Economics. New Brunswick, NJ: Rutgers University Press.
- Munshi, Kaivan 2003. "Networks in the Modern Economy: Mexican Migrants in the U. S. Labor Market", *Quarterly Journal of Economics*, 118(2):549-99
- Oaxaca, Ronald. 1973. "Male-Female Differentials in Urban Labor Markets," *International Economic Review*, 14: 693-709.
- Portes, Alejandro 1987. "The Social Origins of the Cuban Enclave Economy of Miami," *Sociological Perspectives* 30:340-372
- Robert, Stephanie and James S. House, 1996. "SES Differentials in Health by Age and Alternative Indicators of SES," *Journal of Aging and Health*, 8(3): 359-388.
- Sanders, Jimmy M. and Victor Nee, 1996. "Immigrant Self-Employment: The Family as Social Capital and the Value of Human Capital," *American Sociological Review*, 61(2): 231-249.
- Scanlon, Edward, and Deborah Page-Adams. 2001. "Effects of Asset Holding on Neighborhoods, Families, and Children: A Review of Research." Pp. 25–49 in Building Assets: A Report on the Asset-Development and IDA Field, edited by Ray Boshara. Washington, D.C.: Corporation for Enterprise Development.
- Smith, James P. 1995. "Racial and Ethnic Differences in Wealth," *Journal of Human Resources*, 30 (Supplement):S158-S183
- U.S. Board of Governors of the Federal Reserve Board, 2001. *Survey of Consumer Finances*
- U.S. Census Bureau, 2000. *Foreign-Born Population in the United States: Population Characteristics*

**Table 1: Financial Market Participation and Transitions, 1996 – 2000 SIPP Panel**

	All	Native-born	Immigrants	Immigrant/Native
<b>A: Savings Account Ownership</b>				
Own %	52.66%	54.72%	39.72%	0.73
Never Owned %	31.91%	30.18%	42.80%	1.42
Ever Owned %	31.10%	30.77%	33.24%	1.08
Entry %	5.82%	5.92%	5.32%	0.90
Exit %	5.15%	4.86%	7.66%	1.58
Always Owned %	36.99%	39.06%	23.96%	0.61
Observations	356,769	307,894	48,875	
<b>B: Checking Account Ownership</b>				
Own %	34.11%	36.08%	21.74%	0.60
Never Owned %	53.28%	51.33%	65.60%	1.28
Ever Owned %	23.60%	23.93%	21.58%	0.90
Entry %	3.05%	3.15%	2.53%	0.80
Exit %	5.25%	4.98%	8.07%	1.62
Always Owned %	23.11%	24.75%	12.82%	0.52
Observations	356,769	307,894	48,875	
<b>C: Stock Ownership</b>				
Own %	18.42%	20.00%	8.50%	0.42
Never Owned %	73.08%	71.17%	85.12%	1.20
Ever Owned %	15.37%	16.20%	10.13%	0.63
Entry %	1.73%	1.84%	1.10%	0.60
Exit %	7.50%	7.30%	10.61%	1.45
Always Owned %	11.55%	12.63%	4.75%	0.38
Observations	356,769	307,894	48,875	

Note: The sample consists of all MSA residents greater than or equal to the age of 18.

"Own" means that the respondent had a saving account or checking account (interest bearing) during the interview period. "Never Owned" means that the respondent had no saving account or checking account (interest bearing) in all the interview periods. "Ever Owned" means that the respondent had a saving account or checking account (interest bearing) in some of the interview periods, but not all. "Always Owned" means that the respondent had a saving account or checking account (interest bearing) in all the interview periods.

The sum of the percentage of Never Owned, Ever Owned and Always Owned is equal to 1.

Entry is defined as the individual switches from non-ownership to ownership.

Exit is defined as the individual switches from ownership to non-ownership.



**Table 2: Characteristics of the Native-born and Immigrants in the MSA Sample, 1996 – 2000 SIPP Panel**

	All	Natives	Immigrants
Age	45.98 (17.34)	46.18 (17.47)	44.70 (16.41)
Number of Children < 18	0.78 (1.14)	0.72 (1.09)	1.13 (1.36)
Monthly Per Capita Household Income	2116.31 (2764.29)	2195.18 (2810.94)	1619.47 (2391.05)
% Male	45.81%	45.70%	46.46%
% Married	58.45%	57.31%	65.65%
% unemployed or out of the labor force	33.95%	33.48%	36.94%
Race (%)			
White	70.08%	77.15%	25.53%
Black	13.06%	14.20%	5.83%
Hispanic	11.98%	6.97%	43.52%
Asian	4.42%	1.16%	24.93%
Other	0.47%	0.51%	0.20%
Education (%)			
Less than High School	17.86%	15.03%	35.73%
High School Graduate	29.67%	30.48%	24.59%
Some College	29.12%	30.58%	19.95%
College Graduate	15.40%	15.87%	12.42%
Advanced Degree	7.94%	8.04%	7.31%
<u>Immigrant Characteristics</u>			
Years In U.S. (%)			
Less Than 10 Years			37.74%
10 < Duration < 14			17.28%
15 < Duration < 30			16.96%
More Than 30 Years			15.20%
Immigrant Region of Origin (%)			
Central America			32.51%
Asia			20.63%
European			15.11%
Caribbean			7.73%
South America			4.53%
North America			1.62%
Middle East			1.14%
Other			16.73%
Number of Observations	356,769	307,894	48,875

Note: The sample consists of all MSA residents greater than or equal to the age of 18.  
Standard deviations are shown in parentheses ONLY for continuous variables.

**Table 3: Logit Estimates of Financial Market Participation**

	Savings Account		Checking Account	
	(1)	(2)	(1)	(2)
	Coef.	M.E.	Coef.	M.E.
Immigrant	-0.296 *** (0.031)	-0.074	-0.303 *** (0.037)	-0.061
Age	-0.004 (0.003)	-0.001	0.015 *** (0.004)	0.003
Age Squared (x100)	0.014 *** (0.003)	0.004	0.009 ** (0.004)	0.002
Unemployed/Out of Labor Force	-0.293 *** (0.025)	-0.073	0.080 *** (0.028)	0.017
Per Capita HH Income (x100)	0.021 *** (0.001)	0.005	0.021 *** (0.001)	0.004
Per Capita HH Income Squared (x10 <sup>6</sup> )	-0.006 ** (0.0004)	-0.002	-0.005 *** (0.0003)	-0.001
Married	0.873 *** (0.022)	0.215	0.837 *** (0.025)	0.170
Male	-0.299 *** (0.021)	-0.074	-0.268 *** (0.023)	-0.056
Non-White	-0.432 *** (0.026)	-0.108	-0.629 *** (0.032)	-0.121
No of children < 18	-0.082 *** (0.009)	-0.020	-0.080 *** (0.011)	-0.017
High School	0.543 *** (0.030)	0.133	0.776 *** (0.039)	0.172
Some College	0.861 *** (0.031)	0.208	1.177 *** (0.040)	0.264
College	1.037 *** (0.038)	0.241	1.489 *** (0.045)	0.348
Advanced Degree	0.940 *** (0.048)	0.217	1.581 *** (0.053)	0.373
No of Obs	356,769		356,769	
Log-likelihood	-215936.94		-193291	
Pseudo R-squared	0.125		0.156	

Note: The sample consists of all MSA residents greater than or equal to the age of 18.

The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question and is zero otherwise.

Logit model with fixed effects at MSAs level is used and standard errors are corrected for clustering at the individual level.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, schooling dummies and wave dummies. The omitted education category is less than a high school education.

\*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Table 4: Logit Estimates of Financial Market Participation (with duration controls)**

	Savings Account				Checking Account			
	(1) Coef.	M.E.	(2) Coef.	M.E.	(3) Coef.	M.E.	(4) Coef.	M.E.
Immigrant	-0.212 *** (0.034)	-0.053	-0.186 *** (0.059)	-0.047	-0.250 *** (0.040)	-0.051	-0.217 *** (0.067)	-0.044 ***
Recent (after 1990)	-0.501 *** (0.068)	-0.124			-0.354 *** (0.089)	-0.069		
1990-1996			-0.531 *** (0.085)	-0.131			-0.383 *** (0.106)	-0.074
1985-1989			-0.172 * (0.090)	-0.043			-0.186 * (0.111)	-0.038
1980-1984			-0.062 (0.089)	-0.016			-0.142 (0.110)	-0.029
1975-1979			0.192 ** (0.098)	0.048			0.188 (0.119)	0.041
1970-1974			-0.048 (0.110)	-0.012			0.089 (0.126)	0.019
1964-1969			-0.008 (0.122)	-0.002			-0.042 (0.139)	-0.009
1960-1964			-0.036 (0.148)	-0.009			-0.186 (0.166)	-0.038
(Omitted Category: Before 1960)								
No of Obs	356,769		356,769		356,769		356,769	
Log-likelihood	-215760.57		-215718		-193235.03		-193228.2	
Pseudo R-squared	0.156		0.126		0.156		0.156	

Note: The sample consists of all MSA residents greater than or equal to the age of 18.

The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question and is zero otherwise.

Logit model with fixed effects at MSAs level is used and standard errors are corrected for clustering at the individual level.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, schooling dummies and wave dummies. The omitted education category is less than a high school education.

\*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Table 5: Decomposition of Immigrant-Native Differences in Financial Market Participation**

		Participation				
		Saving Acct		Checking Acct		
		Immigrant	Native	Immigrant	Native	
S	Mean:	0.397	0.547	0.217	0.361	
	Gap:	0.150		0.143		
F A		(1)	(2)	(3)	(4)	
U M	Overall Difference:	From ( $X^N - X^I$ )	0.076	0.084	0.083	0.103
L P			50.70%	55.92%	57.76%	71.61%
L L		From ( $\beta^N - \beta^I$ )	0.074	0.066	0.061	0.041
E			49.30%	44.08%	42.24%	28.39%
Contribution to the gap from the following variables:						
Age and Age Square		0.003	0.004	0.005	0.008	
		1.69%	2.63%	3.74%	5.92%	
Per Capita HH Income		0.021	0.019	0.021	0.019	
		14.32%	12.66%	14.96%	13.08%	
Education		0.028	0.036	0.024	0.032	
		18.63%	23.81%	16.88%	22.41%	
Male		0.0007	0.0008	-0.0003	0.0002	
R S		0.46%	0.51%	-0.23%	0.11%	
A A	Marital status	-0.009	-0.017	-0.001	-0.008	
N M		-6.23%	-11.26%	-0.68%	-5.55%	
D P	Non-white	-0.003	0.022	0.007	0.037	
O L		-2.14%	15.00%	4.62%	25.72%	
M E	No of children < 18	0.008	0.006	0.004	0.004	
		5.24%	4.12%	2.87%	2.64%	
Unemploy		0.002	0.002	-0.0002	0.000	
		1.47%	1.52%	-0.16%	-0.33%	
MSA Effects		0.026	0.010	0.023	0.011	
		17.26%	6.92%	15.76%	7.62%	
All variables		0.076	0.084	0.083	0.103	
		50.70%	55.92%	57.76%	71.61%	

Note: The full sample consists of All MSA residents greater than or equal to the age of 18. To keep the native and immigrant samples comparable, some of the MSAs are dropped where MSA fixed effects cannot be estimated separately for the immigrant sample due to a lack of observations.

The random sample includes 10,000 native and 10,000 immigrants randomly drawn from the full sample with replacement.

Column (1) and (3) use the coefficients from the immigrant sample, and Column (2) and (4) use the coefficients from the native sample. See Appendix III for the detailed coefficients.

Logit models with the fixed effects at MSAs level are used and the standard errors are corrected for clustering at the individual level.

The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question, and is zero otherwise.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, and schooling dummies.

The omitted education category is less than high school.

**Table 6: Immigrant Heterogeneity and Financial Market Participation  
(Marginal Effects Only)**

	Ownership					
	(1)Baseline	(2) Race	(3) Legal Status	(4) Greater than high school	(5) Exclude Mexican Immigrants	(6) Hispanics only
<b>A: Savings Acct</b>						
Immigrant	-0.07 ***	-0.13 ***	-0.09 ***	-0.10 ***	-0.05 ***	-0.06 ***
Immi*Non-white		0.16 ***				
Non-white		-0.14 ***				
Immi*Permanent Resident			0.02 *			
Greater Than High School				0.13 ***		
Immi * Greater Than High School				0.03 **		
Number of obs	356,769	356,769	356,769	343,464	42,667	
Log-likelihood	-215531	-215926	-217133	-208910	-22683	
Pseudo R-squared	0.13	0.13	0.12	0.12	0.16	
<b>B: Checking Acct (Interest Bearing)</b>						
Immigrant	-0.06 ***	-0.10 ***	-0.08 ***	-0.11 ***	-0.05 ***	-0.06 ***
Immi*Non-white		0.17 ***				
Non-white		-0.15 ***				
Immi*Permanent Resident			0.03 **			
Greater Than High School				0.15 ***		
Immi * Greater Than High School				0.10 ***		
Number of obs	356,769	356,769	356,769	343,464	42,667	
Log-likelihood	-192971	-193276	-195226	-208910	-22683	
Pseudo R-squared	0.16	0.16	0.15	0.12	0.16	

Note: The sample consists of all MSA residents greater than or equal to the age of 18. The young sample only includes the MSA residents between the age of 18 and 25.

The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question and is zero otherwise.

Logit model is used and standard errors are clustered at the individual level.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, schooling dummies and wave dummies.

The omitted education category is less than high school graduate.

\*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Table 7: Logit Estimates of Financial Market Transitions**

**I. ENTRY INTO ACCOUNT OWNERSHIP**

	(1)		(2)	
	Coef.	M.E.	Coef.	M.E.
<b>A: Savings Acct</b>				
Immigrant	-0.06 (0.04)	-0.003	-0.01 (0.04)	-0.001
Recent (>1990)			-0.21 *** (0.08)	-0.009
No of Obs	145,530		145,530	
Log-likelihood	-30838.41		-30833.90	
Pseudo R-squared	0.041		0.036	
<b>B: Checking Acct</b>				
Immigrant	-0.10 ** (0.05)	-0.002	-0.043 (0.049)	-0.001
Recent (>1990)			-0.309 *** (0.108)	-0.007
No of Obs	204,202		204,202	
Log-likelihood	-26430.93		-26723.14	
Pseudo R-squared	0.051		0.049	

**II. EXITS OUT OF ACCOUNT OWNERSHIP**

	(1)		(2)	
	Coef.	M.E.	Coef.	M.E.
<b>C: Savings Acct</b>				
Immigrant	0.281 *** (0.041)	0.014	0.278 *** (0.040)	0.013
Recent (>1990)			0.019 (0.098)	0.001
No of Obs	164,900		164,900	
Log-likelihood	-32261.92		-32261.89	
Pseudo R-squared	0.036		0.032	
<b>D: Checking Acct</b>				
Immigrant	0.273 *** (0.052)	0.013	0.275 *** (0.055)	0.014
Recent (>1990)			-0.016 (0.129)	-0.001
No of Obs	107299		107299	
Log-likelihood	-21375.79		-21375.79	
Pseudo R-squared	0.035		0.035	

Note: The sample is restricted to individuals over 18 living in MSAs

The dependent variable is equal to one if the individual switches from non-ownership to ownership (Entry) or from ownership to non-ownership (Exit) for savings account or checking account (interest bearing), respectively.

A logit model is used and standard errors are corrected for clustering at the individual level.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, schooling dummies and wave dummies.

The omitted education category is less than high school.

\*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Table 8: The Impact of Location on Financial Market Participation  
(Marginal Effects Only)**

	Ownership		
	(1)	(2)	(3)
<b>A: Savings Acct</b>			
Immigrant	-0.06 ***	-0.03 ***	-0.04 ***
Immi * Ethnic Concentration in MSA	-0.66 ***	-0.72 ***	-0.67 ***
Recent (after 1990)		-0.13 ***	-0.12 ***
Recent * Ethnic Concentration in MSA			-0.53
Number of obs	353083	353083	353083
Log-likelihood	-213488	-213321	-213317
Pseudo R-squared	0.125	0.126	0.126
<b>B: Checking Acct (Interest Bearing)</b>			
Immigrant	-0.04 ***	-0.02 **	-0.02 **
Immi * Ethnic Concentration in MSA	-1.38 ***	-1.42 ***	-1.37 ***
Recent (after 1990)		-0.08 ***	-0.07 ***
Recent * Ethnic Concentration in MSA			-0.74
Number of obs	353,083	353,083	353,083
Log-likelihood	-191138	-191073	-191069
Pseudo R-squared	0.157	0.157	0.157

Note: The sample consists of all MSA residents greater than or equal to the age of 18.

The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question and is zero otherwise.

Logit model is used and standard errors are corrected for clustering at the individual level.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, schooling dummies and wave dummies. The omitted education category is less than high school graduate. \*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Appendix I: Logit Estimates of Financial Market Participation, 1996 – 2000 SIPP Panel  
Additional Indicators of Financial Asset Participation**

	Participation							
	(1)		(2)		(3)		(4)	
	Stock		IRA or Keogh		Mutual Funds		Combined Checking <sup>1</sup>	
	Coef.	M.E.	Coef.	M.E.	Coef.	M.E.	Coef.	M.E.
Immigrant	-0.611 (0.052)	-0.056	-0.436 (0.055)	-0.039	-0.418 (0.058)	-0.031	-0.342 (0.033)	-0.052
Age	0.060 (0.005)	0.006	0.215 (0.006)	0.022	0.057 (0.006)	0.005	0.015 (0.004)	0.002
Age Squared (x100)	-0.036 (0.005)	-0.004	-0.177 (0.006)	-0.018	-0.038 (0.006)	0.000	0.009 (0.004)	0.001
Unemployed/Out of Labor Force	0.241 (0.036)	0.027	0.131 (0.039)	0.013	0.404 (0.041)	0.036	-0.201 (0.027)	-0.029
Per Capita HH Income (x100)	0.024 (0.001)	0.003	0.017 (0.001)	0.002	0.023 (0.001)	0.002	0.025 (0.001)	0.035
Per Capita HH Income Squared (x10 <sup>6</sup> )	-0.005 (0.000)	0.001	-0.004 (0.000)	0.000	-0.005 (0.000)	0.000	-0.006 (0.001)	-0.001
Married	0.758 (0.032)	0.079	0.621 (0.034)	0.061	0.819 (0.036)	0.065	0.980 (0.024)	0.151
Male	-0.224 (0.029)	-0.024	-0.210 (0.031)	-0.021	-0.336 (0.033)	-0.028	-0.372 (0.023)	-0.052
Non-White	-0.770 (0.044)	-0.070	-1.075 (0.051)	-0.085	-0.762 (0.052)	-0.053	-0.805 (0.003)	-0.134
# of children < 18	-0.068 (0.014)	-0.007	-0.153 (0.016)	-0.015	-0.044 (0.015)	-0.004	-0.114 (0.010)	-0.016
High School	1.185 (0.065)	0.155	0.978 (0.066)	0.116	1.242 (0.081)	0.130	0.886 (0.032)	0.109
Some College	1.650 (0.066)	0.233	1.440 (0.066)	0.185	1.737 (0.080)	0.200	1.386 (0.034)	0.163
College	2.155 (0.068)	0.375	2.041 (0.069)	0.334	2.463 (0.082)	0.382	1.769 (0.041)	0.173
Advanced Degree	2.142 (0.074)	0.398	2.166 (0.075)	0.388	2.670 (0.087)	0.465	1.824 (0.054)	0.158
MSA Fixed Effects	YES		YES		YES		YES	
No of Obs	356,769		330,042		330,042		196,787	
Pseudo R-squared	0.180		0.217		0.189		0.207	

<sup>1</sup>The combined checking account indicator measures whether an individual held a checking account (interest bearing or non-interest bearing) during the interview period in question and is zero otherwise. Because non-interest bearing checking is measured less frequently, we have fewer observations for this indicator. Note: The full sample consists of All MSA residents greater than or equal to the age of 18. To keep the native and immigrant samples comparable, some of the MSAs are dropped where MSA fixed effects cannot be estimated separately for the immigrant sample due to a lack of observations. Logit models with the fixed effects at MSAs level are used and the standard errors are corrected for clustering at the individual level. All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, and schooling dummies. The omitted education category is less than high school.



**Appendix II: Top-25 MSAs (Based on Population)**

Metropolitan Statistical Area	% Foreign Born (1990 U.S. census)	Imm Pop 90	% Foreign Born (SIPP 1996 Sample)	Largest Ethnic Concentration			
				Country 1	Ethnic Conc1	Country 2	Ethnic Conc2
New York-Northern New Jersey -Long Island, NY	22.65%	3260551	18.84%	Italy	1.62%	Dominican Republic	1.57%
Los Angeles-Riverside -Orange County, CA	32.98%	2905552	39.03%	Mexico	13.46%	El Salvador	2.00%
Chicago-Gary-Kenosha, IL-IN-WI	15.01%	753332	15.48%	Mexico	4.22%	Poland	1.35%
San Francisco-Oakland -San Jose, CA	23.62%	1164254	29.97%	Mexico	4.37%	Philippines	3.32%
Washington-Baltimore, DC-MD-VA-WV	10.95%	500004	13.00%	El Salvador	0.82%	Korea	0.70%
Philadelphia-Wilmington -Atlantic City, PA-NJ-DE-MD	6.32%	271774	6.46%	Italy	0.57%	Germany	0.51%
Detroit-Ann Arbor-Flint, MI	6.66%	242155	7.09%	Canada	1.14%	Italy	0.50%
Boston-Worcester-Lawrence, MA-NH-ME-CT	12.77%	435377	13.22%	Canada	1.34%	Italy	1.02%
Dallas-Fort Worth, TX	9.41%	265538	14.72%	Mexico	4.28%	Vietnam	0.59%
Houston-Galveston-Brazoria, TX	14.83%	389256	18.46%	Mexico	6.40%	El Salvador	1.09%
Miami-Fort Lauderdale, FL	39.06%	958188	41.98%	Cuba	17.32%	Colombia	2.13%
Seattle-Tacoma-Bremerton, WA	9.31%	203895	14.29%	Canada	1.16%	Philippines	1.02%
Atlanta, GA	4.80%	100422	9.15%	Korea	0.41%	Germany	0.34%
San Diego, CA	19.51%	367263	27.30%	Mexico	8.03%	Philippines	2.75%
Anaheim-Santa Ana -Garden Grove, CA	27.55%	502450	\	Mexico	11.38%	Vietnam	2.68%
Minneapolis-St. Paul, MN	4.37%	78899	6.53%	Laos	0.62%	Canada	0.31%
St. Louis, MO-IL	2.56%	45894	2.92%	Germany	0.31%	Italy	0.16%
Cleveland-Akron, OH	5.94%	105152	3.97%	Yugoslavia	0.78%	Italy	0.46%
Tampa-St. Petersburg -Clearwater, FL	8.37%	137736	8.00%	Canada	1.06%	Cuba	1.01%
Pittsburgh-Beaver Valley, PA	2.95%	47556	3.48%	Italy	0.47%	Germany	0.39%
Phoenix, AZ	8.64%	134719	11.61%	Mexico	3.59%	Canada	0.76%
Denver-Boulder-Greeley, CO	6.31%	93315	8.97%	Mexico	1.61%	Germany	0.57%
Cincinnati-Hamilton, OH-KY-IN	2.34%	29902	1.67%	Germany	0.44%	India	0.17%
Milwaukee-Racine, WI	4.40%	51816	8.79%	Germany	0.70%	Mexico	0.51%
Sacramento-Yolo, CA	12.00%	131261	20.95%	Mexico	2.74%	Philippines	1.00%

Note: The Census sample consists of all MSA residents greater than or equal to the age of 18 in Census 1990 1% Sample.

The SIPP sample consists of all MSA residents greater than or equal to the age of 18 in the wave 2 of 1996-2000 SIPP Panel.

**Appendix III: Logit Estimates of Financial Market Participation, 1996 – 2000 SIPP Panel  
Native and Immigrant Samples**

	Participation			
	Saving Acct		Checking Acct	
	Native	Immigrant	Native	Immigrant
	(1)	(2)	(3)	(4)
Age	-0.007 *	0.016	0.014 ***	0.050 ***
	(0.004)	(0.010)	(0.004)	(0.013)
Age Squared (x100)	0.017 ***	-0.006	0.010 **	-0.031 **
	(0.004)	(0.010)	(0.004)	(0.013)
Unemployed/Out of Labor Force	-0.287 ***	-0.319 ***	0.075 **	0.117
	(0.027)	(0.069)	(0.030)	(0.081)
Per Capita HH Income (x100)	0.020 ***	0.022 ***	0.019 ***	0.031 ***
	(0.001)	(0.003)	(0.001)	(0.002)
Per Capita HH Income Squared (x10 <sup>6</sup> )	-0.006 ***	-0.006 ***	-0.004 ***	-0.008 ***
	(0.0004)	(0.0018)	(0.0003)	(0.0009)
Married	0.879 ***	0.792 ***	0.852 ***	0.646 ***
	(0.024)	(0.062)	(0.027)	(0.079)
Male	-0.298 ***	-0.301 ***	-0.253 ***	-0.409 ***
	(0.023)	(0.057)	(0.025)	(0.069)
Non-White	-0.574 ***	0.069	-0.784 ***	-0.166 **
	(0.030)	(0.059)	(0.037)	(0.072)
# of children < 18	-0.072 ***	-0.113 ***	-0.0666 ***	-0.1186 ***
	(0.011)	(0.023)	(0.012)	(0.029)
High School	0.522 ***	0.464 ***	0.710 ***	0.847 ***
	(0.034)	(0.071)	(0.044)	(0.101)
Some College	0.843 ***	0.763 ***	1.115 ***	1.266 ***
	(0.036)	(0.078)	(0.044)	(0.104)
College	1.046 ***	0.688 ***	1.436 ***	1.397 ***
	(0.042)	(0.093)	(0.050)	(0.115)
Advanced Degree	0.910 ***	0.848 ***	1.500 ***	1.706 ***
	(0.054)	(0.122)	(0.059)	(0.136)
MSA Fixed Effects	YES	YES	YES	YES
No of Obs	302,247	48,822	298,532	48,256
Pseudo R-squared	0.121	0.131	0.148	0.182

Note: The full sample consists of All MSA residents greater than or equal to the age of 18. To keep the native and immigrant samples comparable, MSAs are dropped where MSA fixed effects cannot be estimated separately for the immigrant sample due to a lack of observations.

Logit models with the fixed effects at MSAs level are used and the standard errors are corrected for clustering at the individual level.

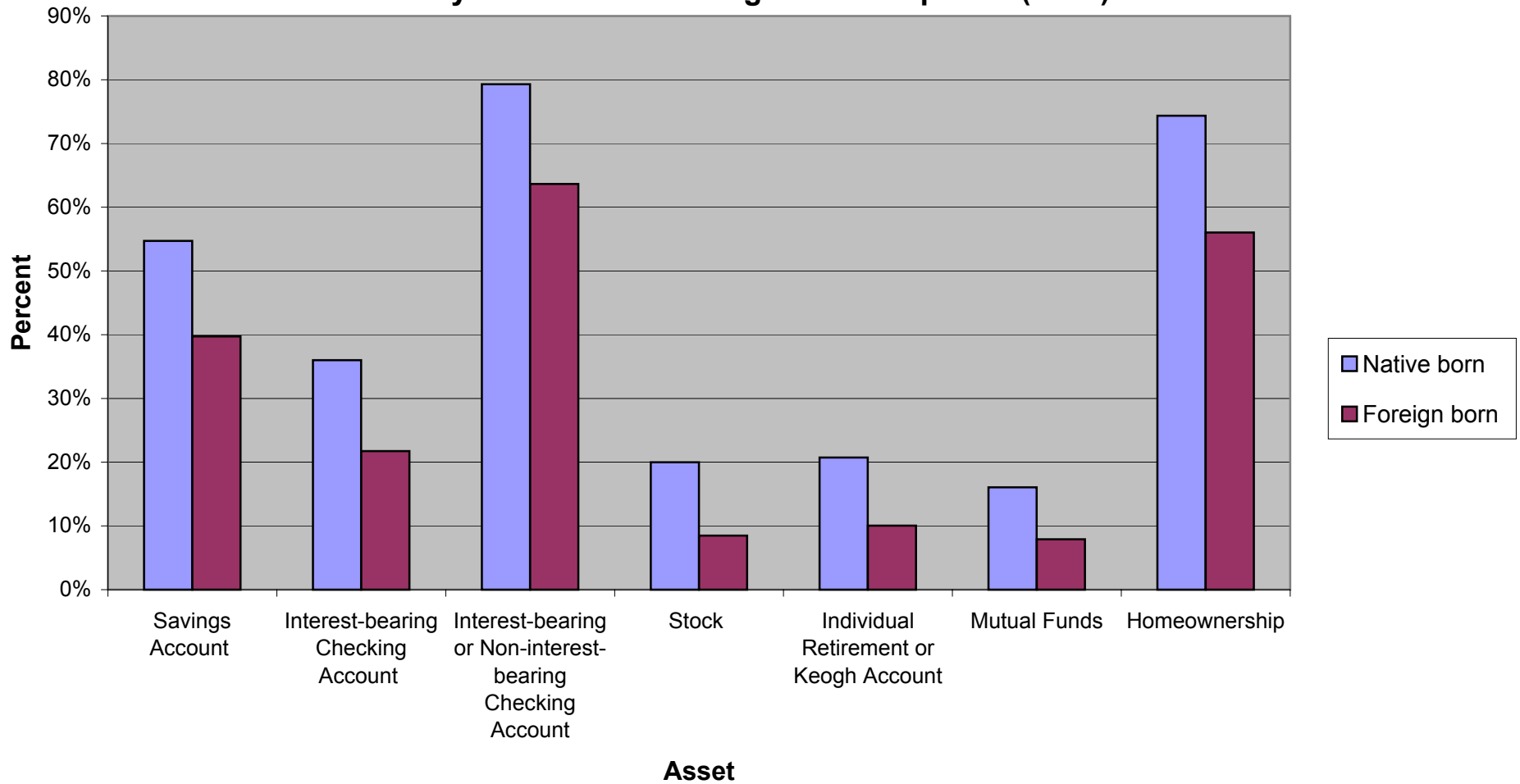
The dependent variable is equal to one if the respondent had a saving account or checking account (interest bearing) during the interview period in question and is zero otherwise.

All regressions include a constant term, age as linear and square terms, per capita income as linear and square terms, marital status, male, non-white, labor force status, number of kids, and schooling dummies.

The omitted education category is less than a high school education.

\*\*\* indicates significance at at least the 1% level, \*\* at at least the 5% level, \* at at least the 10% level.

**Figure 1: Financial Asset Ownership of Native born and Foreign born Individuals**  
**Source: Survey of Income and Program Participation (SIPP) 1996-2000**



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